THE CONTINUOUS-TIME TIME BERNULLI PROCESS

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Abstract: A random Bernoulli process with continuous time and a finite number of states (random events) is proposed. The process is obtained by two mutually complementary methods - directly from the Poisson process with an intensity parameter that depends on time and methods of queuing theory, from a queuing system with two parameters. In the first case, the process was formalized on a probability space with measure, as a measurable function of time. The intensity of the Poisson process was considered as a measure. The Bernoulli process for each fixed time was obtained as a conditional distribution from a suitable Poisson distribution. The parameter of the Poisson distribution was determined from the differential equation, in the formulation of which the approximation of the Bernoulli formula by the Poisson formula was essentially used. In the second method, standard methods of queuing theory were used. A two-parameter queuing model was formulated in which for all customer flows the time between occurrence of neighboring customers was a random value satisfying the exponential law. The model was formalized by a system of differential equations, whose analytical solution represented the continuous-time Bernoulli process. In finding solutions, the method of generating functions was used. It is of interest to derive the Bernoulli process both from the probability space constructed for the Poisson process and from the queuing theory model. The authors believe that the proposed process can be generalized to a wider class of functions than that used in the work, down to measurable ones. The possibilities for the practical application of the continuous-time Bernoulli process will undoubtedly be expanded, since its discrete analog is well known in many fields of science and technology.

Keywords: Poisson, Bernoulli processes; systems of stochastic differential equations; queueing theory; generating function; probability distribution; Markov process

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INTRODUCTION

Mathematical modeling based on the methods of the theory of random processes is very actively used in the study and analysis of the functioning of real objects of varying complexity. Two groups of models can be distinguished - some characterize objects from general positions, in the first approximation and reflect the common, most important trend of development and analysis of real processes, others reveal their hidden essence, help to explore the internal properties of objects. Models of the first group, usually prognostic, are not very parametric, they have good statistics, and analytical solutions, representable by formulas convenient in engineering calculations, or easily computable are preferable for them [1-8]. Models of the second group are multiparameter, heavy, statistics are not always reliable algorithmic methods of investigation are usually effective for them, usually approximate, simulation modeling [9-14], etc. Stochastic modeling is applied to objects, usually associated with randomness, but they are also used in research, even deterministic real objects, using randomization. This often leads to more effective results than traditional mathematical models of the exact sciences

can provide. Many random phenomena and objects representing dynamical systems of a special type are effectively described by methods of queuing theory (OT) [2-8]. In the QT, the queuing system (model) (QS) is first formulated, which, in the continuous case, is formalized by a system of differential equations. The unknown time functions in the systems of equations are the probability distributions usually associated with the Markov random process [2-7, 13-16]. Analytical solutions, although exist, but are found, basically, by approximate methods (in real time in the presence of high-speed computing tools). This is often due to the lack of a special need to look for exact solutions, since private, target problems that are not directly related to mathematics are being studied. In recent years, significant results have been achieved in the development of methods for calculating such systems, but they are reduced to the improvement of numerical algorithms [10]. The model given in this paper, formalized by a system of differential equations, has an analytic solution, which is a random Bernoulli process. When finding solution, the method of generating functions was used [9, 16]. The model is formulated in the terminology of QT. The solution obtained has direct

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practical applications, for example, to the analysis of the efficiency of the functioning of computer systems [2–5, 11, 12], processes arising in chemistry and the chemical industry, biotechnology [5–8, 17], etc.

MATERIALS AND METHODS

Necessary definitions and notations

A probability space is a triple $(\Omega, \mathfrak{T}, P)$, where $\Omega = \{\omega\}$ is the space of elementary random events ω , $\mathfrak{T} - \sigma$ is the algebra of random events, and P is the probability (probabilistic measure, probability distribution) [19–22]. To the fundamental properties of a probability space, we assign a random variable $\xi = \xi(\omega)$ and its mean value – the mathematical expectation $M\xi$ [23].

A random process is a measurable function $\xi(t) = \xi(t, \omega)$ of two variables, one of which is time. As a rule, time fixes the momentary state of the process [23]. For a fixed time *t*, the random process is a random variable [21–23].

The random Bernoulli process in the literature is usually associated with random walks [16, 19, 25, 26]. There are *n* independent tests, in each of which event *A* appears with a constant probability *p* and the event \overline{A} appears with probability q=1-p, then the probability that event A appears exactly *k* times is calculated by the Bernoulli formula

$$P_n(k) = C_n^k p^k q^{n-k}, \ k = 0, 1, 2, \dots n .$$
 (1)

Let $\xi(t)$ be a random variable for any fixed time *t*, characterizing the number of occurrences of the event *A* in *n* independent tests, then the probability $P = \{P_k(t)\}$ is a probability distribution and is defined by the formula

$$P\{\xi = k\} = C_n^k p(t)^k q(t)^{n-k} ,$$

$$\sum_{k=0}^n C_n^k p(t)^k q(t)^{n-k} = 1, \ \forall t \in [0,\infty) .$$
(2)

For each value of the fixed time parameter, let us denote $(\Omega^t, \mathfrak{I}^t, P^t)$, $t \in [0, \infty)$, the probability space of the random Poisson process $\xi(t)$, where $\Omega^t = \{C_k\}$, $\mathfrak{I}^t - \sigma$ is the algebra of events, $P^t = P\{\xi(t) = k\}$. The rejection of the stationarity condition [15, 16, 19, 22] leads to a model of a Poisson process with a variable intensity measure $\lambda(t)$

$$\begin{split} P\{\xi(t) = m\} &= \frac{\lambda(t)^m}{m!} e^{-\lambda(t)}, \ \forall t \in [0,\infty) \\ m &= 0, 1, 2, \dots. \end{split}$$

Let us note that if the intensity of the event stream is given by a nonrandom integrable function $\lambda = \lambda(t)$, and the flow corresponds to the conditions of the absence of aftereffect and ordinariness, it loses the properties of the Poisson process. If the intensity itself is a random function $\lambda(t)$, the process remains Poisson, provided that with $t \to \infty$ if there is a limit, $\lambda(t) \to \lambda > 0$.

Markov process

In the QT terminology of, queue systems (QS) are considered that are described by a random Markov type process with a countable number n of states (starting with zero), postulates of the Poisson process [15, 16, 21, 25, 26] and a set of unknown time functions that form the distribution probabilities.

Suppose that there is a queue system (QS) consisting of a finite set of states $\{C_k\}$, k=0, 1, 2, ..., n. The time of residence of the system in each of the states is random, described by the exponential distribution law. The transition of the system from state to state over time $\Delta t \rightarrow 0$ is carried out abruptly. The assumption of the exponential law of the distribution of the time of residence of the system in each of the states is equivalent to the fulfillment of the condition of absence of aftereffect and ordinariness, and if the parameter of the exponential law is constant, then the stationarity [2, 4, 21, 24]. The fulfillment of these conditions makes it possible to construct a system of linear differential equations for the QS transition from a state to a state determined by a marked state graph, whose structure depends on the object under study [2-9, 14]. Let us denote by $P_k(t,i)$ the probability that at the time $t \in [0, \infty)$, the QS is in the state k, provided that at the initial time t = 0 it was in the state C_i . The random process under consideration is formalized by a system of differential equations that describe the connection between the probabilities that, at time t, there are krequirements in the QS.

Method of generating functions

The method of generating functions makes it possible to convolute a system consisting of an arbitrary number of differential equations for the probability distribution of the system residence in each of the states, to a single partial differential equation of the generating function, and then to solve it.

Suppose we have a probability distribution $\{P_m(t,i)\}$, $\sum_{m=0}^{\infty} P_m(t,i) = 1$, or some other sequence of functions. Let us note each member $\{P_m(t,i)\}$ of this sequence by multiplying it by the complex variable z to the extent that it is equal to the serial number of the sequence member. From the obtained products $z^k \cdot P_k(t,i)$ we will compose the functional series

$$F(z,t) = \sum_{m=0}^{\infty} z^m \cdot P_m(t,i)$$
(3)

Since $|P_m(t,i)| \le 1$ are bounded, the series (3) converges, and, consequently, F(z,t) -is an analytic function in the circle $|z| \le 1$.

The function (3) F(z,t) is called generating for the distribution of the original $\{P_m(t,i)\}$. The original for any *m* is uniquely reconstructed from the formula

$$P_m(t,i) = \frac{1}{m!} \cdot \frac{\partial^m F(0,t)}{\partial z^m}$$
(4)

Next, we need the linearity property and the Laplace image of the derivative of the generating function [4, 5, 22, 24–26].

The linearity property of the generating function makes it possible to reduce the system of differential equations for the unknown distribution $\{P_m(t,i)\}$ to a single partial differential equation of the function F(z,t)

under the initial conditions $F(z,0) = z^{t}$ and $F(0,t) = P_{i}(t,i)$.

RESULTS AND DISCUSSION

Formulation of the model

Let $(\Omega, \mathfrak{T}, P)$ be a probability space in which a Poisson process $\xi = \xi(t)$ is given for each fixed time $t \in [0, \infty)$, where $\Omega = R^+$, I = [0, a), $0 < a < \infty$, $a \in R$; $\mathfrak{I}=B$ are Borel subsets of *I*; *P* is the Poisson distribution.

A Poisson process with a state space defined on a probability space is a measurable function defined on the set of all countable subsets of the Borel set.

Suppose that for each fixed t, I=I(t) -a measurable space of states with finite measure and Borel subsets of its state is defined, then

$$P\{\xi(t) = m\} = V_m(t) = \mu(t)^m \exp(-\mu(t)) / m!$$
(5)

where $\mu(t)$ is the Poisson process intensity parameter, M(t) is a mathematical expectation, $\mu(t) = M(t)$.

According to the definition of density, the intensity $\mu(t)$ can be regarded as a measure of a random process $\xi(t)$ [25, 26], the measure of the space Ω is finite $\mu(t) = \mu(I(t)) < \infty$. It is required to find $P_n(t,k)$ - the probability that for each fixed point of time $t \in [0,\infty)$ and the number of the first points of space Ω , up to *n* inclusive, exactly *k* of them is the event *A* (*t*), then the remaining (*n*-*k*) points belong to the remaining event $\overline{A}(t)$. Taking into account the additive measure property, we have $\mu(A(t)) + \mu(\overline{A}(t)) = \mu(I(t))$. Using the conditional probability formula, we obtain

$$P\{A(t) \land \overline{A}(t) / I(t)\} = V_k(t)V_{n-k}(t) / V_n(t) = \frac{\mu(A(t))^k \exp(-\mu(A(t))\mu(\overline{A}(t))^{n-k} \exp(-\mu(A(t))n!)}{k!(n-k)!\mu(I(t))^n \exp(-I(t))}$$

Or

$$P\{A(t) \wedge \overline{A}(t) / I(t)\} = C_n^k \left[\frac{\mu(A(t))}{\mu(At)) + \mu(\overline{A}(t))} \right]^k \left[\frac{\mu(\overline{A}(t))}{\mu(A(t)) + \mu(\overline{A}(t))} \right]^{n-k}$$
(6)

The probability distribution has the form

$$P_{n}(t,k) = C_{n}^{k} \left[\frac{\mu(A(t))}{\mu(A(t)) + \mu \overline{A}((t))} \right]^{k} \left[\frac{\mu(\overline{A}(t))}{\mu(A(t)) + \mu(\overline{A}(t))} \right]^{n-k}, \quad k = 0, 1, ..., n$$
(7)

It is natural to call the formula (7) the Bernoulli process generated by the Poisson process [15, 25].

Let us introduce the notations

$$p(t) = \frac{\mu(A(t))}{\mu(A(t)) + \mu(\overline{A}(t))}, \quad q(t) = \frac{\mu(A(t))}{\mu(A(t)) + \mu(\overline{A}(t))},$$
$$p(t) + q(t) = 1, \quad t \in [0, \infty).$$
(8)

If the time interval between the occurrences of the neighboring events satisfies the exponential law, then $\mu(A(t)) = np(t)$, $\mu(\overline{A}(t)) = nq(t)$, and $\mu(I(t)) = n$ Setting the initial conditions p(0) = 1, q(0) = 0, taking into account (8), we find

$$p(t) = \frac{\mu}{\lambda + \mu} + \frac{\lambda}{\lambda + \mu} \exp(-(\lambda + \mu)t),$$
$$q(t) = \frac{\lambda}{\lambda + \mu} - \frac{\lambda}{\lambda + \mu} \exp(-(\lambda + \mu)t).$$
(9)

Formulas (9) are solutions of the differential equation [2],

$$\frac{d}{dt}p(t) + \lambda p(t) = \mu$$
(10)

The Bernoulli process can be written in the form

$$P_{n}(t,k) = C_{n}^{k} p(t)^{k} q(t)^{n-k}, \quad p(t) + q(t) = 1,$$

$$t \in [0,\infty), \quad (11)$$

where p(t), q(t) satisfy the formulas (9); initial conditions $P_n(0,n) = 1$, $P_n(0,k) = 0$, $k \neq n$.

Under arbitrary initial conditions, the formulas (11) become significantly more complicated.

Comment

If we go to the limit $t \to \infty$ in formulas (9), putting

$$P_n(k) = \lim_{t \to \infty} P_n(t,k), \quad p = \lim_{t \to \infty} p(t) = \frac{\mu}{\lambda + \mu},$$
$$q = \lim_{t \to \infty} q(t) = \frac{\lambda}{\lambda + \mu},$$

then we get the usual Bernoulli formula, expressed in terms of the intensity parameters λ , μ

$$P_n(k) = C_n^k \left(\frac{\mu}{\lambda + \mu}\right)^k \left(\frac{\lambda}{\lambda + \mu}\right)^{n-k}$$
(12)

Here the sum $\lambda + \mu$ of the parameters is considered

as a measure of the state space of the Poisson process. Let us note that in fact formula (12) is a Bernoulli process for a stationary regime, if we use the terminology of queuing theory (QT) [21].

The Bernoulli process can be obtained in another way, using the methods of queuing theory, within which it becomes clear the appearance of the differential equation (10), from which formulas (9) are derived. For convenience, let us introduce the notations: Let $P_k(t,i) = P_k(t)$, and the initial conditions *i*, denoting the QS state at the initial point of time t = 0, will be discussed separately, if necessary.

Formulation of the model

There is a QS containing *n* customers. Each of them can be in one of two incompatible states. Transition of a customer from one state to another and back satisfies the Poisson postulates with intensities λ and μ , respectively. It is required to find the probability $P_k(t)$ that, at the time $t \in [0, \infty)$ k of the customers are in one state and *n*-k customers are in the other, provided that at the initial moment of time the QS was in the state C_i .

The model is formalized by a system of differential equations [15]

$$\frac{d}{dt}P_{0}(t) = -n \cdot \lambda \cdot P_{0}(t) + \mu \cdot P_{1}(t),$$

$$\frac{d}{dt}P_{k}(t) = -((n-k)\lambda + k\mu) \cdot P_{k}(t) + (n-k+1) \cdot \lambda \cdot P_{k-1}(t) + (k+1) \cdot \mu \cdot P_{k+1}(t), \ 1 \le k < n,$$

$$\frac{d}{dt}P_{n}(t) = -n \cdot \mu \cdot P_{n}(t) + \lambda \cdot P_{n-1}(t),$$
(13)

with initial conditions

$$P_i(0) = 1, P_k(0) = 0, \forall k \neq i, i = 0, 1, ..., n$$
 (14)

The system (13) with the initial conditions (14) is called the Bernoulli process equations system. For the probability distribution $\{P_k(t)\}$, the normalization condition [15] holds,

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$$\sum_{k=0}^{n} P_k(t) = 1,$$
(15)

which is a consequence of the formulation of the model.

Formula (15) is also valid for a countable number of QS states, if $n \to \infty$.

Let us solve the system (13), under the initial conditions (14). Differentiating

$$F(z,t) = \sum_{k=0}^{n} z^{k} \cdot P_{k}(t)$$

with respect to the variable *t* and *z*, we obtain, respectively,

$$\frac{\partial F(z,t)}{\partial t} = \sum_{k=0}^{n} z^{k} \cdot \frac{dP_{k}(t)}{dt}, \quad \frac{\partial F(z,t)}{\partial z} = \sum_{k=1}^{n} k \cdot z^{k-1} \cdot P_{k}(t)$$

Summing the first equation of system (13) with the second, multiplied by z^k and with the last, multiplied by z^n , we obtain

$$\begin{split} \sum_{k=0}^{n} z^{k} \cdot \frac{dP_{k}(t)}{dt} &= -n \cdot \lambda \cdot P_{0}(t) + \mu \cdot P_{1}(t) + \sum_{k=1}^{n-1} z^{k} \cdot \left(-\left[(n-k) \cdot \lambda + k \cdot \mu\right] \cdot P_{k}(t) + (n-k+1) \cdot \lambda \cdot P_{k-1}(t) + \left((k+1) \cdot \mu \cdot P_{k+1}(t)\right) + z^{n} \cdot \left(-n \cdot \mu \cdot P_{N}(t) + \lambda \cdot P_{N-1}(t)\right) = -n \cdot \lambda \cdot P_{0}(t) + \mu \cdot P_{1}(t) - n \cdot \lambda \cdot \sum_{k=1}^{n-1} z^{k} \cdot P_{k}(t) + \left((n-k+1) \cdot \mu \cdot P_{N-1}(t)\right) + z^{n} \cdot \left(-n \cdot \mu \cdot P_{N}(t) + \lambda \cdot P_{N-1}(t)\right) = -n \cdot \lambda \cdot P_{0}(t) + \mu \cdot P_{1}(t) - n \cdot \lambda \cdot \sum_{k=1}^{n-1} z^{k} \cdot P_{k}(t) + z^{n} \cdot \left(-n \cdot \mu \cdot P_{N}(t) + \lambda \cdot P_{N-1}(t)\right) = -n \cdot \lambda \cdot P_{0}(t) + \mu \cdot P_{1}(t) - n \cdot \lambda \cdot \sum_{k=1}^{n-1} z^{k} \cdot P_{k}(t) + z^{n} \cdot \left(-n \cdot \mu \cdot P_{N}(t) + \lambda \cdot P_{N-1}(t)\right) = -n \cdot \lambda \cdot P_{0}(t) + \mu \cdot P_{1}(t) - n \cdot \lambda \cdot \sum_{k=1}^{n-1} z^{k} \cdot P_{k}(t) + z^{n} \cdot \left(-n \cdot \mu \cdot P_{N}(t) + \lambda \cdot P_{N-1}(t)\right) = -n \cdot \lambda \cdot P_{0}(t) + \mu \cdot P_{1}(t) - n \cdot \lambda \cdot \sum_{k=1}^{n-1} z^{k} \cdot P_{k}(t) + z^{n} \cdot \left(-n \cdot \mu \cdot P_{N}(t) + \lambda \cdot P_{N-1}(t)\right) = -n \cdot \lambda \cdot P_{0}(t) + \mu \cdot P_{1}(t) - n \cdot \lambda \cdot \sum_{k=1}^{n-1} z^{k} \cdot P_{k}(t) + z^{n} \cdot \left(-n \cdot \mu \cdot P_{N}(t) + \lambda \cdot P_{N-1}(t)\right) = -n \cdot \lambda \cdot P_{0}(t) + \mu \cdot P_{1}(t) - n \cdot \lambda \cdot \sum_{k=1}^{n-1} z^{k} \cdot P_{k}(t) + z^{n} \cdot \left(-n \cdot \mu \cdot P_{N}(t) + \lambda \cdot P_{N-1}(t)\right) = -n \cdot \lambda \cdot P_{0}(t) + \mu \cdot P_{1}(t) + 2n \cdot P_{N}(t) + 2n \cdot P_{N}(t)$$

$$\begin{split} &+\lambda\sum_{k=1}^{n-1}k\cdot z^{k}\cdot P_{k}(t)-\mu\cdot\sum_{k=1}^{n-1}k\cdot z^{k}\cdot P_{k}(t)+(n+1)\cdot\lambda\cdot\sum_{k=1}^{n-1}z^{k}\cdot P_{k-1}(t)-\lambda\cdot\sum_{k=1}^{n-1}k\cdot z^{k}\cdot P_{k-1}(t)+\\ &+\mu\cdot\sum_{k=1}^{n-1}(k+1)\cdot z^{k}\cdot P_{k+1}(t)-n\cdot\mu\cdot z^{n}\cdot P_{n}(t)+\lambda\cdot z^{n}\cdot P_{n-1}(t)\\ &\frac{dF(z,t)}{dt}=-n\cdot\lambda\cdot P_{0}(t)+\mu\cdot P_{1}(t)-n\cdot\lambda\cdot F(z,t)+n\cdot\lambda\cdot P_{0}(t)+n\cdot\lambda\cdot z^{n}\cdot P_{n}(t)+\lambda\cdot z\cdot\frac{\partial F(z,t)}{\partial z}-\lambda\cdot n\cdot z^{n}\cdot P_{n}(t)-\\ &-\mu\cdot z\cdot\frac{\partial F(z,t)}{\partial t}+\mu\cdot n\cdot z^{N}\cdot P_{n}(t)+(n+1)\cdot\lambda\cdot\sum_{k=0}^{n-2}z^{k+1}\cdot P_{k}(t)-\lambda\cdot\sum_{k=0}^{n-2}(k+1)\cdot z^{k+1}\cdot P_{k}(t)+\mu\cdot\sum_{k=2}^{n}k\cdot z^{k-1}\cdot P_{k}(t)-\\ &-n\cdot\mu\cdot z^{n}\cdot P_{n}(t)+\lambda\cdot z^{n}\cdot P_{n-1}(t)=\mu\cdot P_{1}(t)-n\cdot\lambda\cdot F(z,t)+\lambda\cdot z\cdot\frac{\partial F(z,t)}{\partial z}-\mu\cdot z\cdot\frac{\partial F(z,t)}{\partial t}+\mu\cdot n\cdot z^{n}\cdot P_{N}(t)+\\ &+(n+1)\cdot\lambda\cdot z\cdot F(z,t)-(n+1)\cdot\lambda\cdot z^{n}\cdot P_{n-1}(t)-(n+1)\cdot\lambda\cdot z^{n+1}\cdot P_{n}(t)-\lambda\cdot\sum_{k=2}^{n}k\cdot z^{k+1}\cdot P_{k}(t)-\\ &-\lambda\cdot\sum_{k=2}^{n}z^{k+1}\cdot P_{k}(t)+\mu\cdot\frac{\partial F(z,t)}{\partial z}-\mu\cdot P_{1}(t)-n\cdot\mu\cdot z^{n}\cdot P_{n-1}(t)-(n+1)\cdot\lambda\cdot z^{n+1}\cdot P_{n}(t)-\lambda\cdot z^{2}\cdot\frac{\partial F(z,t)}{\partial z}+\\ &+\lambda(n-1)\cdot z^{n}\cdot P_{n-1}(t)+\lambda\cdot n\cdot z^{n+1}\cdot P_{n}(t)-\lambda\cdot z\cdot F(z,t)+\lambda\cdot z\cdot n\cdot P_{n-1}(t)-(n+1)\cdot\lambda\cdot z^{n+1}\cdot P_{n}(t)-\lambda\cdot z^{2}\cdot\frac{\partial F(z,t)}{\partial z}+\\ &+\lambda\cdot z^{n}\cdot P_{n-1}(t)=-n\cdot\lambda\cdot F(z,t)+\lambda\cdot z\cdot\frac{\partial F(z,t)}{\partial z}-\mu\cdot z\cdot\frac{\partial F(z,t)}{\partial z}+(n+1)\cdot\lambda\cdot z\cdot F(z,t)-\lambda\cdot z^{2}\cdot\frac{\partial F(z,t)}{\partial z}-\\ &-\lambda\cdot z\cdot F(z,t)+\mu\cdot\frac{\partial F(z,t)}{\partial z}-\mu\cdot z\cdot\frac{\partial F(z,t)}{\partial z}-\mu\cdot z\cdot\frac{\partial F(z,t)}{\partial z}+(n+1)\cdot\lambda\cdot z\cdot F(z,t)-\lambda\cdot z^{2}\cdot\frac{\partial F(z,t)}{\partial z}-\\ &-\lambda\cdot z\cdot F(z,t)+\mu\cdot\frac{\partial F(z,t)}{\partial z}-\mu\cdot z\cdot\frac{\partial F(z,t)}{\partial z}-\mu\cdot z\cdot\frac{\partial F(z,t)}{\partial z}+(n+1)\cdot\lambda\cdot z\cdot F(z,t)-\lambda\cdot z^{2}\cdot\frac{\partial F(z,t)}{\partial z}-\\ &-\lambda\cdot z\cdot F(z,t)+\mu\cdot\frac{\partial F(z,t)}{\partial z}-\mu\cdot z\cdot\frac{\partial F(z,t)}{\partial z}-\mu\cdot z\cdot\frac{\partial F(z,t)}{\partial z}+(n+1)\cdot\lambda\cdot z\cdot F(z,t)-\lambda\cdot z^{2}\cdot\frac{\partial F(z,t)}{\partial z}-\\ &-\lambda\cdot z\cdot F(z,t)+\mu\cdot\frac{\partial F(z,t)}{\partial z}=\lambda\cdot n\cdot (z-1)\cdot F(z,t)-(z-1)(\lambda z+\mu)\frac{\partial F(z,t)}{\partial z}-\\ &-\lambda\cdot z\cdot F(z,t)+\mu\cdot\frac{\partial F(z,t)}{\partial z}=\lambda\cdot n\cdot (z-1)\cdot F(z,t)-(z-1)(\lambda z+\mu)\frac{\partial F(z,t)}{\partial z}-\\ &-\lambda\cdot z\cdot F(z,t)+\mu\cdot\frac{\partial F(z,t)}{\partial z}=\lambda\cdot n\cdot (z-1)\cdot F(z,t)-(z-1)(\lambda z+\mu)\frac{\partial F(z,t)}{\partial z}-\\ &-\lambda\cdot z\cdot F(z,t)+\mu\cdot\frac{\partial F(z,t)}{\partial z}=\lambda\cdot n\cdot (z-1)\cdot F(z,t)-(z-1)(\lambda z+\mu)\frac{\partial F(z,t)}{\partial z}-\\ &-\lambda\cdot z\cdot F(z,t)+\mu\cdot \frac{\partial F(z,t)}{\partial z}=\lambda\cdot n\cdot (z-1)\cdot F(z,t)-(z-1)(\lambda z+\mu)\frac{\partial F(z,t)}{\partial z}-\\ &-\lambda\cdot z\cdot F(z,t)+\mu\cdot \frac{\partial F(z,t)}{\partial z}-\\ &-\lambda\cdot z\cdot F(z,t)+\lambda\cdot z\cdot$$

Thus, under the assumptions made, the system of equations (13) is reduced to the following partial differential equation [15–21]:

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$$\frac{\partial F(z,t)}{\partial t} + (z-1)(\lambda z + \mu)\frac{\partial F(z,t)}{\partial z} = N\lambda(z-1)F(z,t)$$
(16)

The general solution of equation (16) can be represented as an arbitrary differentiable function v = g(u), where $u(F, z, t) = c_1$, $v(F, z, t) = c_2$, - are the equations of characteristics. Let us solve the equation (16). For the characteristics we have the system:

$$\frac{dt}{1} = \frac{dz}{(z-1)(\lambda z + \mu)} = \frac{dF}{n \cdot \lambda \cdot (z-1) \cdot F},$$

where F = F(z,t), from which we choose the equations:

$$\frac{dt}{1} = \frac{dz}{(z-1)(\lambda z + \mu)},$$
$$\frac{dz}{(z-1)(\lambda z + \mu)} = \frac{dF}{n \cdot \lambda \cdot (z-1) \cdot F}$$

The solution is found by the standard method, dividing the variables:

$$\frac{\lambda z + \mu}{1 - z} \cdot e^{(\lambda + \mu)t} = c_1, \quad F \cdot (\lambda \cdot z + \mu)^{-n} = c_2.$$

Let us write the solution of equation (16) in the form

$$F \cdot (\lambda \cdot z + \mu)^{-n} = g\left(\frac{\lambda z + \mu}{1 - z} \cdot e^{(\lambda + \mu)t}\right), \quad (17)$$

where g=g(z, t) is an arbitrary continuous and differentiable function.

Further we have

$$F = (\lambda \cdot z + \mu)^n \cdot g\left(\frac{\lambda z + \mu}{1 - z} \cdot e^{(\lambda + \mu)t}\right)$$

Noting that for t = 0, $F(z,0) = z^{i}$, which follows from (14), then

$$z^{i} = (\lambda \cdot z + \mu)^{n} \cdot g\left(\frac{\lambda z + \mu}{1 - z}\right)$$

Let us find the function g, let

$$y = \frac{\lambda z + \mu}{1 - z},$$

then

$$z = \frac{y - \mu}{\lambda + y}$$

For arbitrary *t*, the argument of *g* is (17)

Therefore,

we obtain

In the new notations, we obtain

$$\left(\frac{y-\mu}{\lambda+y}\right)^{l} \cdot \left(y \cdot \frac{\lambda+\mu}{\lambda+y}\right)^{-n} = g(y)$$

or

$$g(y) = \left(\frac{\lambda z + \mu}{1 - z} \cdot e^{(\lambda + \mu)t}\right)^{-n} \cdot \left(\frac{\lambda z + \mu}{1 - z} \cdot e^{(\lambda + \mu)t} - \mu\right)^{i} \cdot (\lambda + \mu)^{-n} \cdot \left(\frac{\lambda z + \mu}{1 - z} \cdot e^{(\lambda + \mu)t} + \lambda\right)^{n-i} = \left(\lambda z + \mu\right)^{-n} \cdot e^{-n(\lambda + \mu)t} \cdot \left((\lambda z + \mu) \cdot e^{(\lambda + \mu)t} - \mu \cdot (1 - z)\right)^{i} \cdot (\lambda + \mu)^{-n} \cdot \left((\lambda z + \mu) \cdot e^{(\lambda + \mu)t} + \lambda \cdot (1 - z)\right)^{n-i}.$$

Taking into account (13), we have:

g(

$$F(z,t) = (\lambda + \mu)^{-N} \cdot \left(\mu \cdot (1 - e^{-(\lambda + \mu)t}) + z \cdot (\lambda + \mu \cdot e^{-(\lambda + \mu)t})\right)^{t} \cdot \left(\mu + \lambda \cdot e^{-(\lambda + \mu)t} + z \cdot \lambda \cdot (1 - e^{-(\lambda + \mu)t})\right)^{N-t} \dots (18)$$

Let us introduce the notations

$$r(t) = \frac{\lambda}{\lambda + \mu} \cdot \left(1 - e^{-(\lambda + \mu)t}\right),$$

$$s(t) = \frac{1}{\lambda + \mu} \cdot \left(\mu + \lambda \cdot e^{-(\lambda + \mu)t}\right),$$

$$(19)$$

$$\varphi(t) = \frac{1}{\lambda + \mu} \cdot \left(\lambda + \mu \cdot e^{-(\lambda + \mu)t}\right),$$

$$\psi(t) = \frac{\mu}{\lambda + \mu} \cdot \left(1 - e^{-(\lambda + \mu)t}\right),$$

$$(20)$$

where s(t) = 1 - r(t), $\psi(t) = 1 - \phi(t)$.

Then (14), taking into account (18), (19), can be written in the following form

$$F(z,t) = \left(s(t) + z \cdot r(t)\right)^{n-i} \cdot \left(\psi(t) + z \cdot \varphi(t)\right)^{i} .$$
(21)

Let us apply the formula (2). Let us note that the function F(z,t) is the product of two functions. Let us assume that

$$A(z) = \left(s(t) + z \cdot r(t)\right)^{n-i}, \ B(z) = \left(\psi(t) + z \cdot \varphi(t)\right)^{i} ,$$

then to find
$$P_k(t)$$
, from formula (2), we can apply the Leibniz formula [8] to the product of the functions (17)

 $\frac{\lambda z + \mu}{1 - z} \cdot e^{(\lambda + \mu)t}$

this

value

for

у,

substituting

$$\frac{d^k}{dz^k} \left(A (z) \cdot B (z) \right) = \sum_{j=0}^k \binom{k}{j} \cdot A^{(k-j)}(z) \cdot B^{(j)}(z)$$

Substituting in it the values of the corresponding derivatives of the functions A(z), B(z) we obtain

$$P_{k}(t) = \sum_{j=0}^{k} {i \choose j} \cdot {n-i \choose k-j} \cdot r^{k-j}(t) \cdot \cdot s^{n-i-k+j}(t) \cdot \varphi^{i}(t) \cdot \psi^{i-j}(t)$$
(22)

Taking into account (19), (20), we finally write down

$$P_{0}(t) = \left(\frac{\mu}{\lambda + \mu} + \frac{\lambda}{\lambda + \mu} \cdot e^{-(\lambda + \mu)t}\right)^{n-i} \cdot \left(\frac{\mu}{\lambda + \mu} - \frac{\mu}{\lambda + \mu} \cdot e^{-(\lambda + \mu)t}\right)^{i}, \quad (23)$$

for k > 0

$$P_{k}(t) = \sum_{j=0}^{k} {\binom{i}{j}} \cdot {\binom{n-i}{k-j}} \cdot \left[\frac{\lambda}{\lambda+\mu} + \frac{\mu}{\lambda+\mu} \cdot e^{-(\lambda+\mu)t}\right]^{j} \cdot \left[\frac{\lambda}{\lambda+\mu} - \frac{\lambda}{\lambda+\mu} \cdot e^{-(\lambda+\mu)t}\right]^{k-j} \cdot \left[\frac{\mu}{\lambda+\mu} - \frac{\mu}{\lambda+\mu} \cdot e^{-(\lambda+\mu)t}\right]^{n-k-i+j} \cdot \left[\frac{\mu}{\lambda+\mu} \cdot e^{-(\lambda+\mu)t}\right]^{n-k-i+j} \cdot \left[\frac{\mu}{\lambda+\mu}$$

For zero initial conditions i=0, for each fixed time, t^{t} , we obtain formula (11), taking into account (9). In the stationary mode of the QS functioning, taking into account equality

$$\sum_{j=0}^{k} {i \choose j} \cdot {n-i \choose k-j} = {n \choose k},$$
(25)

regardless of the initial conditions, we obtain the formula (12) already known to us. 84

The system of equations (13) can be regarded as a formalization of the Bernoulli process with two parameters, considered as intensity measures, with a discrete number of states and a continuous time [22]. The introduction of a measure admits an extension of the class of functions up to measurable [25, 26].

Let us obtain a useful result from a model that has important practical applications [2–8]. We will assume that the customers come from an inexhaustible source of intensity λ , and are served with the intensity μ for the same discipline of the queue and Poisson postulates. Then, assuming $n \to \infty$, we obtain the system of equations from the system (13)

$$\begin{cases} \frac{d}{dt} P_0(t) = -\lambda \cdot P_0(t) + \mu \cdot P_1(t), \\ \frac{d}{dt} P_k(t) = -(\lambda + k\mu) \cdot P_k(t) + (k-1) \cdot \lambda \cdot P_{k-1}(t) + (k+1) \cdot \mu \cdot P_{k+1}(t), \end{cases}$$
(26)

Let us introduce the generating function (1), $F(z,t) = \sum_{k=0}^{\infty} z^k \cdot P_k(t)$. at zero initial conditions $P_0(0) = 1$, $P_k(0) = 0$,

 $F(z,0) = 1, k \neq 0.$

Applying it to the system (26), we obtain a partial differential equation¹

$$\frac{\partial F(z,t)}{\partial t} + \mu(z-1)\frac{\partial F(z,t)}{\partial z} = \lambda(z-1)F(z,t)$$
(27)

By analogy with finding the solution of the partial differential equation (16), the solution of this equation under the initial conditions, i=0, has the form

$$F(z,t) = \exp(\lambda / \mu(z-1)(1-\exp(-\mu t)))$$

For the probability distribution we obtain

$$P_{k}(t) = \frac{(M(t)^{k})}{k!} \exp(-M(t)) M(t) = \exp(\lambda / \mu(1 - \exp(-\mu t)))$$

This is the Poisson distribution with a variable parameter $\lambda(t) = M(t)$.

Comment. It was shown in [7, 8] that for relatively arbitrary initial conditions, we can obtain from equation (27) the following system of ordinary differential equations for the mathematical expectation and variance

$$\begin{bmatrix}
\frac{d}{dt}M_{i}(t) + \lambda \cdot M_{i}(t) = \mu, \\
\frac{d}{dt}\left[D_{i}(t) + M_{i}^{2}(t) - M_{i}(t)\right] + 2\lambda\left[D_{i}(t) + M_{i}^{2}(t) - M_{i}(t)\right] = 2\mu \cdot M_{i}(t),$$
(28)

Under the initial conditions

$$M_i(0) = i D_i(0) = 0$$
(29)

The solution of (28), taking (29) into account, has the form

$$\begin{cases} M_{i}(t) = \frac{\mu}{\lambda} + \left(i - \frac{\mu}{\lambda}\right) \cdot e^{-\lambda \cdot t}, \\ D_{i}(t) = (1 - e^{-\lambda \cdot t}) \cdot \left(\frac{\mu}{\lambda} + i \cdot e^{-\lambda \cdot t}\right). \end{cases}$$
(30)

All solutions in a stationary mode: Probability distribution, mathematical expectation and variance are known characteristics of the Poisson process

$$p_k = (\lambda / \mu)^k / k! (\exp(-\lambda / \mu) M = \lambda / \mu D = \lambda / \mu$$

and the probability distribution and moments (30), for each fixed value of time, and zero initial conditions, are shown to be a Poisson process with two parameters formalized by the system of equations of Corollary 1.

¹The first equation (27) and its solution were obtained by Palm S. in 1947. The generating functions were introduced by L. Euler in the middle of the 18th century, considering numerical series.

CONCLUSION

The paper presents a random Bernoulli process with continuous time. A probability space is defined on which a real random variable is given that is a measurable function associated with the theory of a measure that has a Poisson distribution. A space with a countable set of states of a Poisson process is defined for each fixed point of the time axis. As a measure of a countable number of subsets of a σ -algebra, a finite measure of intensity, depending on time, is adopted. Convergence is guaranteed by the Poisson process existence theorem [18]. Consideration of a finite

number of the first *n* states of the Poisson process led to the creation of the Bernoulli process with two parameters, *n* is the measure of the state space and the measure of the intensity is the parameter $\mu(t)$, where $\mu(t) = n - \lambda(t)$. The methods of queuing theory formulated by the QS formalized by a system of differential equations, the solution which is the continuous-time Bernoulli process.

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